#### 4 Regularization and variable selection with LASSO and CVaR penalty

#### 4.1. Introduction

Besides interpretability, an important concern in variable selection is the predictive accuracy of selected model. Accurate out-of-sample forecast can be difficult to get if the data set has outliers. When outliers are in the in-sample set, selecting the relevant variables of the model becomes a difficult problem, with the risk of including variables that are not in the "true" model, but can explain the outliers, called in this thesis "spurious variables". In that case, inclusion of "spurious variables" improves in-sample forecast, but out-of-sample performance will not be satisfactory. In order to avoid this "spurious" selection, we propose an extension of LASSO methodology robust to outliers. The idea is to add a Conditional Value at Risk (CVaR) of "out-of-sample" errors term to the LASSO  $\ell_1$ -penalty, as explained in next section.

CVaR, or expected shortfall, is a risk measure widely used in the recent literature. Known to have better properties than VaR (Value at Risk), CVaR can capture events deep in the tail of the distribution (catastrophic events). Generally speaking, CVaR is the conditional expected value of losses above the VaR. For more details see Alexander and Baptista (2004).

Rockafellar and Uryasev (2000) proposed a technique that calculates the VaR and optimizes CVaR simultaneously, formulating the CVaR as a linear optimization problem. The formulation for the CVaR of a random variable R, proposed by the authors, is presented in (31):

$$CVaR_{\alpha}(\boldsymbol{R}) = \min_{(z,\delta_k)} \left\{ z + \frac{1}{K(1-\alpha)} \sum_{k=1}^{K} \delta_k \right\}$$
(31)

subject to:

$$\begin{split} \mathbf{z} &\geq \mathbf{0}, & \forall \ k = 1, \dots, K \\ \delta_k &\geq R_k - \mathbf{z}, & \forall \ k = 1, \dots, K \end{split}$$

where z is the VaR<sub> $\alpha$ </sub> and  $\alpha$  is the CVaR confidence level. Common values for  $\alpha$  are 0.95 and 0.99.

In this chapter, we use the risk measure CVaR, as presented in (31), in a variable selection problem.

#### 4.2. LASSO-CVaR

We propose a penalized least square criterion based on the LASSO  $\ell_1$ -penalty and CVaR (Conditional Value at Risk) of out-of-sample regression errors. We call this approach LASSO-CVaR. The idea is to select variables controlling the model out-of-sample performance, therefore, we add a CVaR of out-of-sample errors in the penalty term. We believe that LASSO-CVaR method will be capable to identify outliers, not selecting "spurious variables" that would increase the out-of-sample error.

The  $\text{CVaR}_{\alpha}$  term in the penalty will minimize the expected value of the largest  $(1 - \alpha)$ % errors out-of-sample, in other words, the  $\text{CVaR}_{\alpha}$  will be the conditional expected value of the out-of-sample errors larger than a  $\text{VaR}_{\alpha}$  value z. In this work, we chose to set  $\alpha$  at 75%.

Consider model estimation and variable selection in a linear regression framework. Suppose that  $\mathbf{y} = (y_1, ..., y_{T_{in}})'$  is the response vector, and  $\mathbf{x}_j = (x_{j1}, ..., x_{jT_{in}})'$ , with j = 1, ..., p, are the predictor variables. Suppose also  $\varepsilon_t = y_t - \boldsymbol{\beta}^T \boldsymbol{x}_t$  and  $\boldsymbol{\varepsilon} = (\varepsilon_{T_{in}+1}, ..., \varepsilon_{T_{out}})'$  is the "out-of-sample" errors vector. The LASSO-CVaR estimator is given by (32):

$$\widehat{\boldsymbol{\beta}}^{LASSO-CVaR} = \arg\min_{\beta} \left\| \mathbf{y} - \sum_{j=1}^{p} \mathbf{x}_{j} \beta_{j} \right\|^{2} + \lambda \sum_{j=1}^{p} |\beta_{j}| + \gamma CVaR_{\alpha}(|\boldsymbol{\varepsilon}|)$$
(32)

where  $\|.\|$  denotes the standard  $\ell_2$ -norm in the "in-sample" set,  $\lambda$  is the

nonnegative regularization parameter of the  $\ell_1$ -penalty (or LASSO penalty) and  $\gamma$  is a nonnegative parameter that gives the weight of CVaR term in the penalty term. More clearly, (32) can be represented as:

$$\widehat{\boldsymbol{\beta}}^{LASSO-CVaR} = \arg \min_{\boldsymbol{\beta}} \sum_{t=1}^{T_{in}} \left( y_t - \sum_{j=1}^p x_{jt} \beta_j \right)^2 + \lambda \sum_{j=1}^p |\beta_j|$$

$$+ \gamma \text{CVaR}_{\alpha} \{ |\varepsilon_t| \}_{t=T_{in}+1}^{T_{out}}$$
(33)

where  $t = 1, ..., T_{in}, T_{in} + 1, ..., T_{out}$ .  $T_{in}$  and  $T_{out}$  represent the number of "insample" and "out-of-sample" observations, respectively. In this context, we are using the quotation marks ("") for "in-sample" and "out-of-sample" because we want to emphasize these are subsets of observations within the original in-sample set. We will always have the true out-of-sample set unknown at the moment of model estimation. In equation (33) the sets  $\{1, ..., T_{in}\}$  and  $\{T_{in} + 1, ..., T_{out}\}$  can be seen as the training and validation sets in neural networks context.

Using the CVaR formulation proposed in Rockafellar and Uryasev (2000) in eq. (31), and taking the vector of explicative variables  $x_t = (x_{1t}, ..., x_{pt})$ , we can rewrite (33) as a quadratic optimization problem as follows:

$$\min_{(\boldsymbol{\beta},\delta_{t},\boldsymbol{z})} \sum_{t=1}^{T_{in}} (\boldsymbol{y}_{t} - \boldsymbol{\beta}^{\mathrm{T}} \boldsymbol{x}_{t})^{2} + \lambda |\boldsymbol{\beta}| + \gamma \left( \boldsymbol{z} + \frac{1}{T_{out}(1-\alpha)} \sum_{t=T_{in}+1}^{T_{out}} \delta_{t} \right)$$
set to:
$$(34)$$

subject to:

$$\delta_t \ge 0, \qquad \forall t = T_{in} + 1, \dots T_{out}$$
  
$$\delta_t \ge |y_t - \boldsymbol{\beta}^{\mathrm{T}} \boldsymbol{x}_t| - z, \quad \forall t = T_{in} + 1, \dots T_{out}$$

# 4.2.1. Theoretical results for the parameter $\lambda$

Let  $\mathbf{y} = (y_1, ..., y_T)'$  be the response vector and  $\mathbf{X}$  be the  $T \times p$  matrix of predictors, with row  $\mathbf{x}_t = (x_{1t}, ..., x_{pt})$ , and column  $\mathbf{x}_j = (x_{j1}, ..., x_{jT})'$ . Consider the original LASSO problem in (35) which is similar to (1):

$$\widehat{\boldsymbol{\beta}}^{LASSO} = \arg\min_{\boldsymbol{\beta}} \frac{1}{2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_{2}^{2} + \lambda \|\boldsymbol{\beta}\|_{1}$$
(35)

where  $\|.\|_2$  denotes the  $\ell_2$ -norm and  $\|.\|_1$  denotes the  $\ell_1$ -norm.

Equation (35) represents a convex but not differentiable function due to the  $\ell_1$ -norm. Effectively the function  $f(\boldsymbol{\beta}) = |\boldsymbol{\beta}|$  is not differentiable in  $\boldsymbol{\beta} = 0$ . The optimality condition in this case is that the subgradient of function  $f(\boldsymbol{\beta})$ includes the point  $\boldsymbol{\beta} = 0$ . Therefore, the optimal solution for (35) must satisfies

$$-\mathbf{X}^{\mathrm{T}}(\mathbf{y} - \mathbf{X}\boldsymbol{\beta}) + \lambda \nu = 0$$
(36)

where  $v_i$  is the *j*th component of the subgradient of  $\|\boldsymbol{\beta}\|_1$ , such that

$$\nu_{j} \in \begin{cases} \{+1\} & \text{if } \beta_{j} > 0\\ \{-1\} & \text{if } \beta_{j} < 0\\ [-1,1] & \text{if } \beta_{j} = 0 \end{cases}$$
(37)

We can write (36) as:

$$\mathbf{X}^{\mathrm{T}}(\mathbf{y} - \mathbf{X}\boldsymbol{\beta}) = \lambda \boldsymbol{\nu} \tag{38}$$

Setting  $\beta = 0$ , by the Karush-Kuhn-Tucker (KKT) conditions, we have

$$\mathbf{X}^{\mathrm{T}}\mathbf{y} \in [-\lambda, \lambda] \tag{39}$$

So, for  $\beta = 0$  the optimal condition for (35) is

$$-\lambda \le \mathbf{X}^{\mathrm{T}} \mathbf{y} \le \lambda \tag{40}$$

or

$$\lambda \ge \max[\mathbf{X}^{\mathrm{T}}\mathbf{y}] \tag{41}$$

Let  $\lambda_{\text{max}}$  be the smallest tuning parameter value for which all coefficients in the solution are zero ( $\beta = 0$ ), for the original LASSO in (35) we have

$$\lambda_{\max} = \max[\mathbf{X}^{\mathrm{T}}\mathbf{y}] \tag{42}$$

This is the maximum value in the sequence of  $\lambda$ 's used in LASSO and adaLASSO estimation. Section 2.3.2 discuss the issue of selecting the parameter  $\lambda$ .

Analogously to (36), if we derive the LASSO-CVaR in (34) and define  $\varepsilon_t = y_t - \boldsymbol{\beta}^T \boldsymbol{x}_t$ , for  $t = T_{in} + 1, ..., T_{out}$ , we have

$$-\mathbf{X}^{\mathrm{T}}(\mathbf{y} - \mathbf{X}\boldsymbol{\beta}) + \lambda \nu + \frac{\gamma}{T_{out}(1-\alpha)} \sum_{t \in \{\mathrm{L}(\varepsilon_t)\}} \eta_t = 0$$
(43)

where  $v_j$  is defined by (37), {L( $\varepsilon_t$ )} is the set of the  $(1 - \alpha)$ % largest  $\varepsilon_t$  and  $\eta_t$  is the subderivate of  $\varepsilon_t$  defined in (44).

$$\eta_t \in \begin{cases} \{-\boldsymbol{x}_t\} & \text{if } \varepsilon_t > 0\\ \{\boldsymbol{x}_t\} & \text{if } \varepsilon_t < 0\\ [-\boldsymbol{x}_t, \boldsymbol{x}_t] & \text{if } \varepsilon_t = 0 \end{cases}$$
(44)

When  $\beta = 0$  we have  $\varepsilon_t = y_t$ . To simplify, we assume that  $y_t \neq 0$ , so from (43), (37) and (44), we have

$$-\lambda \le \mathbf{X}^{\mathrm{T}}\mathbf{y} + \frac{\gamma}{T_{out}(1-\alpha)} \sum_{t \in \{\mathrm{L}(y_t)\}} \mathbf{x}_t \operatorname{sgn}(y_t) \le \lambda$$
(45)

where  $sgn(y_t)$  is the sign function of  $y_t$  and  $t = T_{in} + 1, ..., T_{out}$ .

Let  $\overline{x}_{out}$  be the average of  $(x_t \operatorname{sgn}(y_t))$  with  $t \in \{L(y_t)\}$ , where  $\{L(y_t)\}$  is the set of the  $(1 - \alpha)$ % largest  $y_t$ , as presented in (46):

$$\overline{\boldsymbol{x}}_{out} = \frac{1}{T_{out}(1-\alpha)} \sum_{t \in \{\mathcal{L}(\boldsymbol{y}_t)\}} \boldsymbol{x}_t \operatorname{sgn}(\boldsymbol{y}_t)$$
(46)

it follows that

$$\mathfrak{A} \ge \max[\mathbf{X}^{\mathrm{T}}\mathbf{y} + \gamma \overline{\mathbf{x}}_{out}] \tag{47}$$

Finally,  $\lambda_{max}$  of the LASSO-CVaR is given by eq. (48)

$$\lambda_{\max}^{LASSO-CVaR} = \max[\mathbf{X}^{\mathrm{T}}\mathbf{y} + \gamma \overline{\mathbf{x}}_{out}]$$
(48)

#### 4.3. Simulation

The goal of this simulation exercise is to test the robustness of LASSO-CVaR proposed in this chapter. Therefore, we need to generate a data set with outliers.

Consider the following data generating process (DGP1):

$$y_{t} = \sum_{k=1}^{q} \beta_{k} x_{k,t} + \left(\beta_{q+1} x_{q+1,t} * I_{F_{in}}(t)\right) + 0.5 \varepsilon_{t}, \qquad \varepsilon_{t} \sim \text{IN}[0,1],$$

$$I_{F_{in}}(t) = \begin{cases} 1, & \text{if } t \in F_{in} \\ 0, & \text{if } t \notin F_{in} \end{cases}$$

$$x_{t} = \boldsymbol{v}_{t}, \qquad \boldsymbol{v}_{t} \sim \text{IN}_{q}[0, \boldsymbol{I}_{q}] \quad \text{for } t = 1, \dots, T$$

$$(49)$$

where  $\boldsymbol{\beta}$  is a vector of ones of size q;  $\beta_{q+1} = 5$ ;  $\boldsymbol{x}_t$  is a vector of q relevant variables; and  $F_{in}$  is a set of 5% of the  $T_{in}$  observations. Observations in  $F_{in}$ , chosen randomly, suffer the effect of  $\beta_{q+1}$ , called "fake coefficient".

The term "fake coefficient" is used to give the idea of a coefficient that is not actually relevant in the true model, but still affects the response in some observations. The variable  $\mathbf{x}_{q+1}$  is irrelevant in 95% of the "in-sample" set, i.e.,  $\beta_{q+1} = 5$  in 5% of the "in-sample" observations, and  $\beta_{q+1} = 0$  for the rest of the sample. In other words, we are "contaminating" the "in-sample" data with outliers.

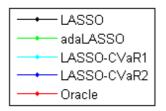
The value for  $\beta_{q+1}$  was fixed in order to produce "big outliers" and increase the variance of the in-sample  $y_t$ . There was no clear base for the choice of this value, and we could have chosen differently, or test model selection methods for different values of  $\beta_{q+1}$ , likewise the choice of the number of "fake coefficients". These are issues for future study and simulations.

In order to test variable selection and out-of-sample performance of the approach proposed in this chapter, we compare the LASSO-CVaR to the original LASSO and adaLASSO for linear regressions as implemented in Chapter 2. We also compare the methods to the oracle approach that provides a best-case scenario by assuming the true model was known. The true model does not include the "fake coefficient".

Similarly to simulation in Chapter 2, the comparison takes into account variable selection statistics, properties of estimators and forecast accuracy. We also compare statistics related to the "fake coefficient" and its selection rate.

LASSO-CVaR solves the quadratic optimization problem in (34) using interior point methods. Due to computational time limitations we used the regularization parameter  $\lambda$  chosen by LASSO, using BIC criterion. In the future, we shall discuss how to build a sequence of  $\lambda$ 's from the  $\lambda_{max}^{LASSO-CVaR}$  in (48) and selection criteria. In this simulation exercise we tested the LASSO-CVaR with  $\gamma = 0.25$  and  $\gamma = 0.5$ . The confidence level  $\alpha$  of the CVaR was set at 0.75, and the "in-sample"  $(T_{in})$  and "out-of-sample"  $(T_{out})$  sets are 80% and 20% of the total *T* observations. We consider a total of 100 out-of-sample observations. For instance, if T = 300, we will have  $T_{in} = 240$ ,  $T_{out} = 60$ , and 100 more out-of-sample (real out-of-sample) observations. As in Section 2.4, we simulate T = 50, 100, 300, 500 observations of DGP1 (49) for different combinations of candidate (*n*) and relevant (*q*) variables. We consider n = 100, 300 and q = 5, 10, 15, 20.

Figures 15-18 illustrate the distribution of the bias for the Oracle, LASSO, adaLASSO, LASSO-CVaR1 (with  $\gamma = 0.25$ ) and LASSO-CVaR2 (with  $\gamma = 0.5$ ) estimators for the parameter  $\beta_1$ , chosen arbitrarily, over 1000 Monte Carlo replications, for different sample sizes, number of candidate variables and number of relevant variables. Figures 19-22 illustrate the distribution of the bias for the "fake coefficient" estimates. In this case it does not make sense to talk about oracle estimator. Color lines of each model are shown in the color legend:



Color Legend for Figures 15-22

As in Section 2.4, from the plots in Figures 15-18, we notice that bias and variance decrease when *T* increases. As expected, the adaLASSO estimator is the closest to the Oracle, but the LASSO-CVaR estimator presented smaller bias and variance than the LASSO, especially when *T* increases. LASSO-CVaR2 ( $\gamma = 0.5$ ) presents smaller bias than LASSO-CVaR1 ( $\gamma = 0.25$ ), which is logical as the  $\gamma = 0.5$  forces the CVaR term to be smaller than  $\gamma = 0.25$ , which should reduces de bias and variance of the estimators.

When comparing distributions for the "fake parameter" in Figures 19-22, we notice that when T = 300 and T = 500, the LASSO-CVaR presents much better results than the others. For T = 50 and T = 100, all models present similar bias, but LASSO-CVaR presents smaller variance.

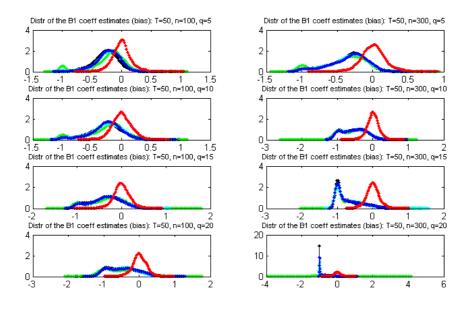


FIGURE 15. Distribution of the bias for the Oracle (red), LASSO (black), adaLASSO (green), LASSO-CVaR1 (cyan) and LASSO-CVaR2 (blue) estimators for the parameter  $\beta_1$  over 1000 Monte Carlo replications. Different combinations of candidate (*n*) and relevant (*q*) variables. The sample size equals 50 observations.

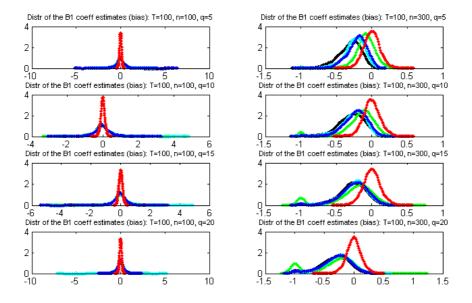


FIGURE 16. Distribution of the bias for the Oracle (red), LASSO (black), adaLASSO (green), LASSO-CVaR1 (cyan) and LASSO-CVaR2 (blue) estimators for the parameter  $\beta_1$  over 1000 Monte Carlo replications. Different combinations of candidate (*n*) and relevant (*q*) variables. The sample size equals 100 observations.

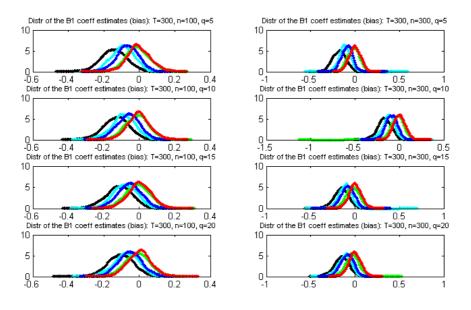


FIGURE 17. Distribution of the bias for the Oracle (red), LASSO (black), adaLASSO (green), LASSO-CVaR1 (cyan) and LASSO-CVaR2 (blue) estimators for the parameter  $\beta_1$  over 1000 Monte Carlo replications. Different combinations of candidate (*n*) and relevant (*q*) variables. The sample size equals 300 observations.

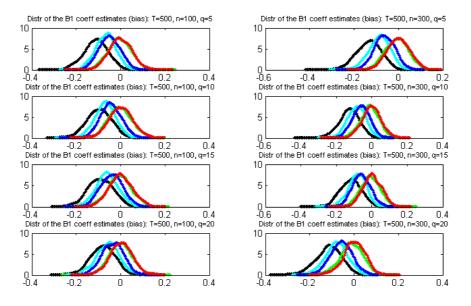


FIGURE 18. Distribution of the bias for the Oracle (red), LASSO (black), adaLASSO (green), LASSO-CVaR1 (cyan) and LASSO-CVaR2 (blue) estimators for the parameter  $\beta_1$  over 1000 Monte Carlo replications. Different combinations of candidate (*n*) and relevant (*q*) variables. The sample size equals 500 observations.

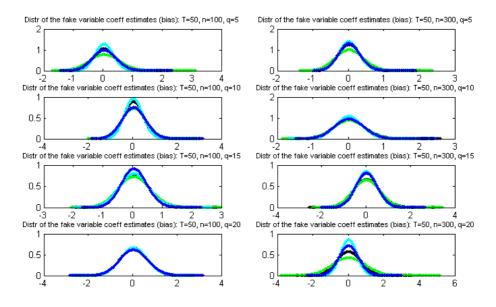


FIGURE 19. Distribution of the bias for the Oracle (red), LASSO (black), adaLASSO (green), LASSO-CVaR1 (cyan) and LASSO-CVaR2 (blue) estimators for the "fake parameter" over 1000 Monte Carlo replications. Different combinations of candidate (n) and relevant (q) variables. The sample size equals 50 observations.

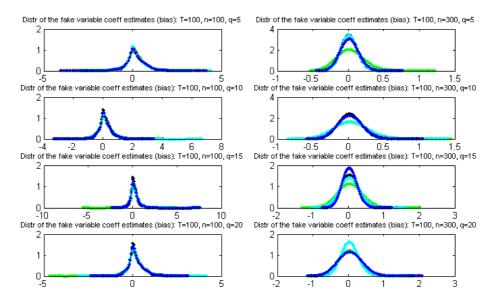


FIGURE 20. Distribution of the bias for the Oracle (red), LASSO (black), adaLASSO (green), LASSO-CVaR1 (cyan) and LASSO-CVaR2 (blue) estimators for the "fake parameter" over 1000 Monte Carlo replications. Different combinations of candidate (n) and relevant (q) variables. The sample size equals 100 observations.

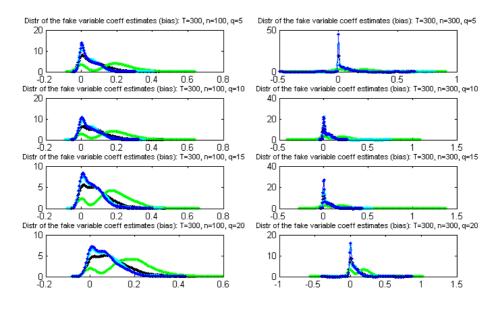


FIGURE 21. Distribution of the bias for the Oracle (red), LASSO (black), adaLASSO (green), LASSO-CVaR1 (cyan) and LASSO-CVaR2 (blue) estimators for the "fake parameter" over 1000 Monte Carlo replications. Different combinations of candidate (n) and relevant (q) variables. The sample size equals 300 observations.

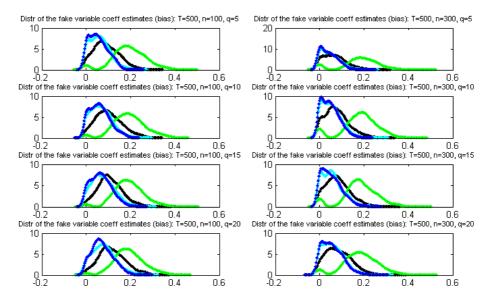


FIGURE 22. Distribution of the bias for the Oracle (red), LASSO (black), adaLASSO (green), LASSO-CVaR1 (cyan) and LASSO-CVaR2 (blue) estimators for the "fake parameter" over 1000 Monte Carlo replications. Different combinations of candidate (n) and relevant (q) variables. The sample size equals 500 observations.

Tables 45-48 present variable selection results for LASSO, adaLASSO, LASSO-CVaR1 and LASSO-CVaR2, following the format and statistics of Tables 2-5 in Section 2.4.1. Comparing Tables 45 and 46 to Tables 4 and 5, we notice that, in some scenarios, with the presence of outliers, LASSO and adaLASSO included more irrelevant variables (Panel (f)). We attribute this

negative result to the "fake coefficient". The LASSO and adaLASSO methods will select irrelevant variables, including the "fake variable", trying to explain the effect caused by the "fake coefficient". We believed the LASSO-CVaR would be able to exclude more irrelevant variables, however Table 47 and 48 show that LASSO-CVaR present worst results than LASSO and adaLASSO concerning variable selection. We notice that LASSO-CVaR1 includes less irrelevant variables (Panel (f)) than LASSO-CVaR2. This can be explained by the increase of parameter  $\gamma$  that forces LASSO-CVaR to decrease the CVaR of the "out-of-sample" errors. In order to do so, the method selects more variables. A possible way of improving this result is reducing  $\gamma$ . This is an important issue for future research.

Table 49 presents the selection rates for the "fake coefficient". Figures 19-22 show that the LASSO-CVaR minimizes the bias and variance of the "fake coefficient", however, Table 49 shows the rate of its inclusion is significantly high, even if lower than for LASSO and adaLASSO when T = 500. This combined analysis gives an indication that, differently from LASSO and adaLASSO, even when LASSO-CVaR selects the "fake coefficient", the method's estimative for this coefficient is close to zero, like in the true model.

#### TABLE 45. MODEL SELECTION: DESCRIPTIVE STATISTICS LASSO

The table reports for each different sample size, several statistics concerning model selection. Panel (a) presents the fraction of replications where the correct model has been selected. Panel (b) shows the fraction of replications where the relevant variables are all included. Panel (c) presents the fraction of relevant variables included. Panel (d) shows the fraction of irrelevant variables excluded. Panel (e) presents the average number of included variables. Panel (f) shows the average number of included irrelevant regressors.

	LASSO											
		=50		100		300		500				
q\n	5	10	5	10	5	10	5	10				
				l (a): Correc	t Sparsity Pa	attern						
5	0.011	0.002	0.012	0.064	0.076	0.101	0.027	0.039				
10	0	0	0	0.001	0.013	0.015	0.004	0.009				
15	0	0	0	0	0.003	0.007	0.001	0.003				
20	0	0	0	0	0.001	0	0.001	0.001				
Panel (b): True Model Included												
5	0.948	0.871	0.983	0.995	1	1	1	1				
10	0.948	0.295	0.985	0.995	1	1	1	1				
10	0.810	0.295	0.975	0.907	1	1	1	1				
20	0.45	0.01	0.932	0.943	1	1	1	1				
20	0.004	0	0.945	0.834	1	T	T	T				
Panel (c): Fraction of Relevant Variables Included												
5	0.985	0.966	0.997	0.998	1	1	1	1				
10	0.965	0.835	0.998	0.991	1	1	1	1				
15	0.914	0.639	0.997	0.983	1	1	1	1				
20	0.813	0.508	0.997	0.965	1	1	1	1				
		D		ation of luna		hlee Fuelud	ما					
5	0.841	0.899	0.123	0.982	0.967	ibles Exclude 0.985	<u>20</u> 0.968	0.989				
5 10	0.841	0.899	0.123	0.982	0.987	0.985	0.968	0.989				
10 15	0.773	0.888	0.109	0.936	0.938	0.976	0.943	0.980				
20	0.749	0.883	0.081	0.881	0.905	0.964	0.917	0.971				
20	0.755	0.001	0.005	0.850	0.809	0.949	0.890	0.902				
			Panel (e	): Number o	of Included \	/ariables						
5	20.072	34.612	88.319	10.365	8.128	9.489	7.998	8.128				
10	30.113	40.751	90.174	28.336	15.591	16.868	15.168	15.718				
15	35.041	42.785	93.083	48.702	23.06	25.344	22.083	23.126				
20	37.439	43.499	93.116	61.289	30.494	34.248	28.839	30.681				
		P	anel (f): Nur	nber of Incl	uded Irrelev	ant Variable	<u>es</u>					
5	15.149	29.783	83.336	5.376	3.128	4.489	2.998	3.128				
10	20.466	32.405	80.195	18.422	5.591	6.868	5.168	5.718				
15	21.333	33.204	78.131	33.963	8.060	10.344	7.083	8.126				
20	21.170	33.344	73.181	41.990	10.494	14.248	8.839	10.681				

#### TABLE 46. MODEL SELECTION: DESCRIPTIVE STATISTICS adaLASSO

The table reports for each different sample size, several statistics concerning model selection. Panel (a) presents the fraction of replications where the correct model has been selected. Panel (b) shows the fraction of replications where the relevant variables are all included. Panel (c) presents the fraction of relevant variables included. Panel (d) shows the fraction of irrelevant variables excluded. Panel (e) presents the average number of included variables. Panel (f) shows the average number of included irrelevant regressors.

			<del>C33013.</del>	daLASSO							
	<u>T</u> =	=50	<u>T</u> =	100		300	<u>T</u> =	500			
q\n	5	10	5	10	5	10	5	10			
			Pane	(a): Correct	t Sparsity Pa	attern					
5	0.003	0	0.004	0.016	0.023	0.05	0.006	0.013			
10	0	0	0	0	0.002	0.002	0.001	0.003			
15	0	0	0	0	0	0.001	0.002	0.001			
20	0	0	0	0	0	0	0	0			
Panel (b): True Model Included											
5	0.878	0.817	0.957	0.978	1	1	1	1			
10	0.794	0.282	0.937	0.923	1	1	1	1			
15	0.444	0.012	0.917	0.848	1	1	1	1			
20	0.069	0.001	0.915	0.718	1	1	1	1			
		_									
_				ction of Rel							
5	0.921	0.911	0.991	0.990	1	1	1	1			
10	0.922	0.782	0.992	0.961	1	1	1	1			
15	0.863	0.602	0.994	0.902	1	1	1	1			
20	0.742	0.483	0.994	0.826	1	1	1	1			
		Da	nel (d): Era	ction of Irre	lovant Varia	bles Exclud	ad				
5	0.838	0.895	0.105	0.977	0.952	0.967	0.953	0.984			
10	0.759	0.895	0.105	0.924	0.932	0.960	0.933	0.972			
15	0.743	0.882	0.000	0.873	0.870	0.948	0.888	0.961			
20	0.743	0.881	0.075	0.875	0.870	0.948	0.853	0.901			
20	0.747	0.001	0.000	0.050	0.025	0.522	0.000	0.547			
			Panel (e	): Number o	f Included \	/ariables					
5	20.023	35.596	89.952	11.671	9.561	14.696	9.452	9.677			
10	30.892	42.268	91.971	31.616	17.888	21.727	17.164	18.055			
15	34.825	42.779	93.714	49.612	26.068	29.798	24.515	26.218			
20	35.088	43.109	93.492	56.207	33.961	41.748	31.721	34.734			
		Pa	anel (f): Nur	nber of Incl	uded Irrelev	ant Variable	es				
5	15.417	31.042	84.997	6.721	4.561	9.696	4.452	4.677			
10	21.674	34.451	82.047	22.003	7.888	11.727	7.164	8.055			
15	21.873	33.752	78.809	36.086	11.068	14.798	9.515	11.218			
20	20.257	33.448	73.603	39.683	13.961	21.748	11.721	14.734			

### TABLE 47. MODEL SELECTION: DESCRIPTIVE STATISTICS LASSO-CVaR1

The table reports for each different sample size, several statistics concerning model selection. Panel (a) presents the fraction of replications where the correct model has been selected. Panel (b) shows the fraction of replications where the relevant variables are all included. Panel (c) presents the fraction of relevant variables included. Panel (d) shows the fraction of irrelevant variables excluded. Panel (e) presents the average number of included variables. Panel (f) shows the average number of included irrelevant regressors.

	LASSO-CVaR1											
		=50		100	<u>T</u> =	300	<u>T</u> =	500				
q\n	5	10	5	10	5	10	5	10				
			Pane	(a): Correc	t Sparsity Pa	attern						
5	0	0	0	0	0	0	0	0				
10	0	0	0	0	0	0	0	0				
15	0	0	0	0	0	0	0	0				
20	0	0	0	0	0	0	0	0				
Panel (b): True Model Included												
5	0.958	0.894	0.975	0.988	1	1	1	1				
10	0.848	0.336	0.95	0.977	1	1	1	1				
15	0.481	0.013	0.934	0.952	1	1	1	1				
20	0.097	0	0.921	0.863	1	1	1	1				
		_										
_					evant Varial							
5	0.987	0.972	0.995	0.996	1	1	1	1				
10	0.971	0.850	0.995	0.992	1	1	1	1				
15	0.919	0.656	0.995	0.986	1	1	1	1				
20	0.835	0.538	0.996	0.970	1	1	1	1				
		Da	nol (d): Era	tion of Irro	levant Varia	bloc Evolud	od					
5	0.761	0.881	0.125	0.942	0.875	0.931	0.875	0.942				
10	0.701	0.878	0.097	0.891	0.875	0.919	0.875	0.942				
15	0.706	0.875	0.075	0.831	0.320	0.915	0.791	0.923				
20	0.700	0.875	0.069	0.838	0.736	0.885	0.754	0.891				
20	0.701	0.074	0.005	0.010	0.750	0.005	0.754	0.051				
			Panel (e	): Number o	of Included \	/ariables						
5	27.596	39.840	88.070	22.161	16.905	25.278	16.905	22.041				
10	35.368	43.865	91.218	41.495	25.692	33.440	24.835	32.345				
15	38.770	45.386	93.587	60.961	33.845	43.165	32.745	41.268				
20	40.632	45.996	94.403	70.963	41.099	52.280	39.690	50.480				
		Pa	anel (f): Nur	nber of Incl	uded Irrelev	<u>ant V</u> ariable	es					
5	22.663	34.978	83.095	17.181	11.905	20.278	11.905	17.041				
10	25.657	35.362	81.268	31.577	15.692	23.44	14.835	22.345				
15	24.983	35.547	78.655	46.170	18.845	28.165	17.745	26.268				
20	23.940	35.244	74.485	51.555	21.099	32.280	19.690	30.480				

### TABLE 48. MODEL SELECTION: DESCRIPTIVE STATISTICS LASSO-CVaR2

The table reports for each different sample size, several statistics concerning model selection. Panel (a) presents the fraction of replications where the correct model has been selected. Panel (b) shows the fraction of replications where the relevant variables are all included. Panel (c) presents the fraction of relevant variables included. Panel (d) shows the fraction of irrelevant variables excluded. Panel (e) presents the average number of included variables. Panel (f) shows the average number of included irrelevant regressors.

number of f	LASSO-CVaR2												
	<u></u>	=50	<u>T</u> =	100	<u>T</u> =	300	<u>T</u> =	500					
q\n	5	10	5	10	5	10	5	10					
			Panel	l (a): Correc	t Sparsity Pa	attern							
5	0	0	0	0	0	0	0	0					
10	0	0	0	0	0	0	0	0					
15	0	0	0	0	0	0	0	0					
20	0	0	0	0	0	0	0	0					
Panel (b): True Model Included													
5	0.972	0.881	0.978	0.998	1	1	1	1					
10	0.852	0.332	0.963	0.979	1	1	1	1					
15	0.517	0.012	0.936	0.957	1	1	1	1					
20	0.110	0	0.915	0.856	1	1	1	1					
		Pa	anel (c): Fra	ction of Rel	evant Varial	bles Include	d						
5	0.993	0.970	0.996	0.999	1	1	1	1					
10	0.974	0.846	0.996	0.996	1	1	1	1					
15	0.932	0.655	0.995	0.985	1	1	1	1					
20	0.843	0.523	0.995	0.970	1	1	1	1					
		Pa	inel (d): Frad	ction of Irre	levant Varia	bles Exclud	ed						
5	0.747	0.882	0.103	0.900	0.715	0.841	0.704	0.818					
10	0.711	0.878	0.092	0.867	0.663	0.819	0.648	0.793					
15	0.702	0.875	0.069	0.831	0.613	0.799	0.607	0.775					
20	0.697	0.873	0.074	0.814	0.568	0.779	0.571	0.758					
			Panel (e)	): Number o	f Included \	/ariables							
5	29.036	39.644	90.229	34.455	32.110	51.785	33.159	58.814					
10	35.716	43.783	91.727	48.488	40.373	62.420	41.689	70.135					
15	39.289	45.517	94.033	63.063	47.859	72.427	48.389	79.246					
20	41.074	45.902	93.970	71.470	54.571	81.752	54.308	87.729					
		De	anal (f): Nur	nhor of Incl	udad Irrala	ant Variabl	05						
5	24.073	34.793	anel (f): Nur 85.251	29.459	27.110	46.785	28.159	53.814					
5 10	24.073 25.981	34.793 35.326	85.251 81.764	29.459 38.532	30.373	40.785 52.42	28.159 31.689	60.135					
10	25.981	35.526 35.689	79.102	38.532 48.285	30.373	52.42 57.427	33.389	64.246					
20	25.315 24.207	35.446	79.102	48.285 52.066	32.859 34.571	57.427 61.752	33.389	64.246 67.729					
20	24.207	55.440	74.003	52.000	74.771	01.752	34.300	07.723					

#### TABLE 49. SELECTION OF "FAKE COEFFICIENT": SELECTION RATE DGP1

replications.								
	<u></u>	=50	<u>T</u> =	100	<u>T=</u>	300	<u>T=</u>	500
q/n	5	10	5	10	5	10	5	10
				LAS	<u>SSO</u>			
5	28%	18%	89%	12%	72%	55%	93%	87%
10	31%	17%	91%	21%	81%	61%	96%	90%
15	29%	17%	92%	30%	83%	67%	97%	91%
20	29%	14%	92%	25%	86%	65%	98%	91%
				<u>adaL</u>	<u>ASSO</u>			
5	25%	18%	90%	12%	77%	58%	95%	90%
10	30%	17%	93%	20%	84%	64%	97%	92%
15	29%	17%	93%	25%	86%	68%	98%	92%
20	28%	13%	92%	21%	89%	67%	98%	91%
				LASSO	-CVaR1			
5	32%	19%	90%	20%	70%	58%	87%	81%
10	34%	16%	92%	27%	75%	62%	91%	85%
15	34%	14%	93%	32%	80%	66%	93%	85%
20	35%	15%	94%	29%	82%	66%	93%	86%
				LASSO	-CVaR2			
5	37%	20%	92%	24%	66%	53%	77%	71%
10	36%	17%	92%	27%	68%	58%	82%	74%
15	35%	18%	94%	35%	72%	62%	83%	75%
20	31%	15%	93%	30%	77%	58%	86%	78%

The table reports for each different sample size, number of candidate variables (n) and number of relevant variables (q), the selection rate for the "fake coefficient" over 1000 Monte Carlo replications.

Finally, Table 50 presents the mean squared error (MSE) for out-ofsample forecast for LASSO, adaLASSO, LASSO-CVaR1, LASSO-CVaR2 and oracle models. We consider a total of 100 out-of-sample observations. As expected, all methodologies improved their performance as the sample size increases, and the number of relevant and candidate variables decrease. LASSO and adaLASSO presented larger MSE than in Table 6, and LASSO-CVaR1 and LASSO-CVaR2 presented smaller MSE than the others (closer to the Oracle). This result can indicates that LASSO and adaLASSO are less capable than LASSO-CVaR of identifying outliers in the "in-sample" data. As expected, when  $\gamma$  increases (LASSO-CVaR2), the MSE out-of-sample decreases.

### TABLE 50. FORECASTING: DESCRIPTIVE STATISTICS DGP1

The table reports for each different sample size, the out-of-sample mean squared error (MSE) for
each model selection technique. $n$ is the number of candidate variables whereas $q$ is the number of
relevant regressors.

relevant reg	<i>T</i> =50			<u></u>		300	<u>T=500</u>		
q/n	5	<u>-50</u> 10	5	100 10	5	10	5	<u>10</u>	
				MSE -	<u>Oracle</u>				
5	0.391	0.401	0.320	0.319	0.272	0.271	0.262	0.262	
10	0.568	0.571	0.392	0.388	0.295	0.295	0.275	0.277	
15	0.840	0.801	0.478	0.479	0.317	0.319	0.291	0.288	
20	1.125	1.180	0.563	0.568	0.340	0.343	0.301	0.305	
					LASSO				
5	1.422	1.997	30.728	0.920	0.393	0.552	0.336	0.364	
10	2.874	6.001	23.204	1.703	0.492	0.643	0.395	0.455	
15	6.104	12.286	23.027	2.734	0.568	0.847	0.446	0.536	
20	11.010	18.573	17.206	4.646	0.630	0.933	0.476	0.619	
				1465					
-	4 0 2 5	2.045	44 2 42		daLASSO	0.014	0.040	0.267	
5	1.925	2.815	41.243	0.722	0.396	0.911	0.349	0.367	
10	3.915	7.291	37.725	1.785	0.458	0.800	0.385	0.416	
15	7.399	13.778	36.909	3.950	0.510	0.776	0.417	0.452	
20	13.546	19.799	33.633	7.114	0.560	0.996	0.439	0.500	
				MSE - LAS	SO_C\/₂R1				
5	1.363	1.917	41.192	0.787	0.331	0.704	0.298	0.312	
10	2.698	5.716	37.364	1.530	0.396	0.677	0.334	0.361	
15	5.855	12.028	33.495	2.540	0.460	0.647	0.363	0.415	
20	10.582	18.174	28.748	4.411	0.506	0.760	0.390	0.466	
	10.002	10.17 1	20.7 10		0.500	0.700	0.000		
				MSE - LAS	SO-CVaR2				
5	1.363	1.966	45.588	0.706	0.339	0.498	0.302	0.326	
10	2.717	5.802	33.811	1.449	0.395	0.498	0.331	0.370	
15	5.718	12.092	32.668	2.547	0.439	0.672	0.357	0.407	
20	10.433	18.384	23.955	4.429	0.484	0.669	0.375	0.453	

Comparing to LASSO and adaLASSO, the LASSO-CVaR method presented vantages concerning forecasting accuracy and disadvantages in variable selection of the true model, when the data presents outliers. One may argue that outliers would hardly be only in the "in-sample" set, and there is an old discussion on how to split the "in-sample" and "out-of-sample" set. With this motivation, we evaluated a new simulation exercise, "contaminating" also the "out-of-sample" data with outliers.

Consider now another data generating process (DGP2):

$$y_{t} = \sum_{k=1}^{q} \beta_{k} x_{k,t} + \left(\beta_{q+1} x_{q+1,t} * \mathbf{I}_{F_{in}}(t)\right) + \left(\beta_{q+1} x_{q+1,t} * \mathbf{I}_{F_{out}}(t)\right) + 0.5 \varepsilon_{t}$$

$$\varepsilon_{t} \sim \mathrm{IN}[0,1]$$

$$\mathbf{I}_{F_{in}}(t) = \begin{cases} 1, & \text{if } t \in F_{in} \\ 0, & \text{if } t \notin F_{in} \end{cases}$$

$$\mathbf{I}_{F_{out}}(t) = \begin{cases} 1, & \text{if } t \in F_{out} \\ 0, & \text{if } t \notin F_{out} \end{cases}$$

$$\mathbf{x}_{t} = \mathbf{v}_{t}, \qquad \mathbf{v}_{t} \sim \mathrm{IN}_{q}[0, \mathbf{I}_{q}] \quad \text{for } t = 1, \dots, T \end{cases}$$

$$(50)$$

where  $\boldsymbol{\beta}$  is a vector of ones of size q,  $\boldsymbol{x}_t$  is a vector of q relevant variables, the "fake coefficient"  $\beta_{q+1} = 5$ ,  $F_{in}$  is the set of 5% of  $T_{in}$  observations and  $F_{out}$  is the set of 5% of  $T_{out}$  observations. Now, the "fake coefficient" is active also in the "out-of-sample" set.

We omitted the plots of the distribution of the bias for the parameter  $\beta_1$ and the "fake coefficient" because they are very similar to Figures 15-22. Variable selection statistics also are very close to the case of DGP1, so tables are omitted as well. However, with the presence of outliers in the "out-of-sample" data, the "fake coefficient" selection rate has increased for all methods as presented in Table 51. When *T* increases the "fake coefficient" selection rate is almost 100% for all methods. This can be explained by the increasing of  $F_{out}$  that increases proportionally with *T*, increasing the number of outliers.

Table 52 reports the MSE for out-of-sample forecast. Comparing to Table 50, we notice that the MSE increased with outliers in the "out-of-sample" set, however LASSO-CVaR still present better results than LASSO and adaLASSO in forecasting accuracy for most of the scenarios.

## TABLE 51. SELECTION OF "FAKE COEFFICIENT": SELECTION RATE DGP2

The table	e reports fo	or eac	ch di	ifferent sau	mple	size	, nur	nber of	candidate va	ariable	es (n) a	and num	ber of
relevant	variables	(q),	the	selection	rate	for	the	"fake	coefficient"	over	1000	Monte	Carlo
replication	ons.												

replications		=50	T=	100	T=	300	<i>T</i> =500		
q/n	5	10	5	10	5	10	5	10	
				LAS	<u>SSO</u>				
5	39%	26%	93%	16%	86%	72%	99%	97%	
10	40%	21%	93%	26%	89%	77%	100%	96%	
15	35%	16%	94%	34%	92%	78%	100%	97%	
20	35%	17%	94%	32%	93%	78%	99%	98%	
					<u>ASSO</u>				
5	32%	24%	93%	14%	90%	75%	99%	98%	
10	37%	19%	93%	23%	92%	80%	100%	96%	
15	32%	15%	94%	26%	94%	79%	99%	97%	
20	31%	16%	94%	23%	95%	79%	100%	97%	
					-CVaR1				
5	42%	30%	92%	22%	82%	67%	95%	89%	
10	44%	23%	93%	31%	86%	74%	97%	91%	
15	44%	19%	96%	37%	88%	75%	97%	94%	
20	37%	16%	94%	34%	90%	77%	98%	95%	
				LASSO	-CVaR2				
5	47%	29%	94%	31%	84%	73%	94%	91%	
10	45%	22%	94%	36%	83%	77%	95%	92%	
15	42%	18%	94%	38%	88%	79%	95%	91%	
20	38%	17%	94%	34%	91%	79%	97%	94%	

### TABLE 52. FORECASTING: DESCRIPTIVE STATISTICS DGP2

The table reports for each different sample size, the out-of-sample mean squared error (MSE) for
each model selection technique. $n$ is the number of candidate variables whereas $q$ is the number of
relevant regressors.

Televalit Teg	-		100	T 000 T 500				
		=50		100		300		500
q/n	5	10	5	10	5	10	5	10
				MSE -	Oracle			
5	0.456	0.454	0.333	0.331	0.277	0.277	0.265	0.264
10	0.694	0.718	0.422	0.422	0.302	0.304	0.281	0.280
15	1.050	1.079	0.520	0.538	0.331	0.331	0.296	0.299
20	1.460	1.503	0.645	0.641	0.360	0.362	0.313	0.311
				<u>MSE -</u>	LASSO			
5	2.056	2.593	40.396	1.048	0.420	0.981	0.363	0.391
10	3.759	6.804	31.166	1.995	0.542	0.794	0.434	0.494
15	7.293	13.003	28.559	3.304	0.624	1.097	0.491	0.598
20	11.793	18.963	24.969	5.733	0.719	1.153	0.540	0.677
				<u>MSE - ac</u>	daLASSO			
5	2.678	3.469	51.710	0.854	0.435	1.388	0.385	0.410
10	5.279	8.042	48.695	2.203	0.510	1.279	0.427	0.456
15	9.431	14.435	44.637	5.019	0.579	1.591	0.460	0.509
20	14.458	20.247	44.142	8.884	0.645	1.293	0.493	0.550
				<u>MSE - LAS</u>	SO-CVaR1			
5	2.283	2.844	58.685	0.914	0.356	0.852	0.312	0.331
10	4.187	6.888	49.706	1.857	0.435	0.691	0.358	0.401
15	7.510	12.836	39.364	3.398	0.517	0.993	0.395	0.470
20	11.951	18.863	37.837	5.794	0.577	1.089	0.430	0.533
				MSE - LAS	SO-CVaR2			
5	2.409	2.732	62.583	0.960	0.365	1.011	0.320	0.344
10	4.184	6.840	46.086	2.052	0.439	0.680	0.357	0.398
15	7.464	12.935	40.923	3.503	0.506	1.010	0.392	0.462
20	11.767	18.694	33.983	5.861	0.582	0.993	0.429	0.515

#### 4.4. Conclusion

In this chapter we introduce an extension of LASSO with a second regularization term. For this, we use the risk measure, widely used in optimization literature, CVaR (Conditional Value at Risk).

Analyzing the results in Section 4.3, we conclude that the LASSO-CVaR has presented good features when the focus is the predictive accuracy of selected models, showing better results in out-of-sample forecasting. However, our goal is the specification of the model selecting the relevant variables of the true model simultaneously. We want to be able to interpret the model and to understand the effects of the explanatory variables on the response. In that matter, LASSO-CVaR

presented worst results than the original LASSO.

Nevertheless, these are the first results for LASSO-CVaR, and many details have to be adjusted. We identified a promising field in the blend of variable selection methods based in shrinkage and the risk measure CVaR. The CVaR term showed to be useful when the data set has outliers and made the LASSO-CVaR a robust method of estimation and variable selection.

However, there are several remain questions to address. In future research we will propose an adaptive version of the LASSO-CVaR (adaLASSO-CVaR). It is important to study more carefully the parameter  $\gamma$ , its importance and selection criteria. Likewise, it would be interesting to estimate several LASSO-CVaR models using a sequence of  $\lambda$ 's, as in the original LASSO.